

APPENDIX A

Derivation of the Weibull related functions

For convenience of derivation, I assume the location parameter $t_0 = 0$.

The percentile of the 2-parameter Weibull distribution t_p than can be directly determined from the CDF_W:

$$F_W(t) = 1 - e^{\left(-\left(\frac{t}{\eta}\right)^\beta\right)}. \quad (\text{A-1})$$

Rearranging equation (A-1), we obtain:

$$\left(t_p\right)_W = \eta \left(\ln \frac{1}{1-p}\right)^{\frac{1}{\beta}}. \quad (\text{A-2})$$

The median is the 50th percentile, hence substituting 0.5 into p yields:

$$\text{Median}_W = \eta (\ln 2)^{\frac{1}{\beta}}. \quad (\text{A-3})$$

The mode is derived by setting $f'_W(t) = 0$, and solving for the maximum values of t.

$$f'_W(t) = \frac{\beta}{\eta} \left(\frac{t}{\eta}\right)^{\beta-1} e^{\left(-\left(\frac{t}{\eta}\right)^\beta\right)}, \quad (\text{A-4})$$

$$\begin{aligned}
 f'_W(t) &= \frac{\beta}{\eta}(\beta - 1) \left[\left(\frac{t}{\eta} \right)^{\beta - 2} \right] \left(\frac{d \left(\frac{t}{\eta} \right)}{d t} \right) e^{-\left(\frac{t}{\eta} \right)^\beta} \\
 &\quad + \frac{\beta}{\eta} \left[\left(\frac{t}{\eta} \right)^{\beta - 1} \right] e^{-\left(\frac{t}{\eta} \right)^\beta} \left[\frac{d \left(- \left(\frac{t}{\eta} \right)^\beta \right)}{d t} \right] \\
 &= \frac{\beta}{\eta^2} e^{-\left(\frac{t}{\eta} \right)^\beta} \left[(\beta - 1) \left(\frac{t}{\eta} \right)^{\beta - 2} - \left(\frac{t}{\eta} \right)^{\beta - 1} \beta \left(\frac{t}{\eta} \right)^{\beta - 1} \right] \\
 &= \frac{\beta}{\eta^2} e^{-\left(\frac{t}{\eta} \right)^\beta} \left(\frac{t}{\eta} \right)^{\beta - 2} \left[(\beta - 1) - \left(\frac{t}{\eta} \right)^\beta \beta \right]
 \end{aligned}$$

For convenience of derivation, we set $q = \frac{t}{\eta}$ then

$$f'(t) = \frac{\beta}{\eta^2} e^{-q^\beta} q^{\beta - 2} (-\beta q^\beta + \beta - 1) \quad (\text{A-5})$$

Let the last term of equation (A-5) = 0 to solve for the maximum value i.e. the mode.

$$-\beta q^\beta + \beta - 1 = 0$$

$$\frac{\beta - 1}{\beta} = \left(\frac{t}{\eta} \right)^\beta$$

Finally the value of the mode can be obtained if $\beta \geq 1$

$$t_{\text{mode}} = \eta \left(\frac{\beta - 1}{\beta} \right)^{\frac{1}{\beta}} = \text{Mode} \quad (\text{A - 6})$$

In order to determine the inflection point, the second derivative is needed.

$$f''(t) = \frac{\beta}{\eta^3} e^{-q^\beta} q^{\beta-3} \left[(\beta q^\beta)^2 - (3\beta - 3)\beta q^\beta + (\beta - 2)(\beta - 1) \right] \quad (\text{A - 7})$$

Equating the last term of equation (A-7) = 0, the inflection point can be obtained by solving for the root of the equation.

$$(\beta q^\beta)^2 - (3\beta - 3)\beta q^\beta + (\beta - 2)(\beta - 1) = 0$$

$$q^\beta = \frac{3(\beta - 1) \pm \sqrt{(5\beta - 1)(\beta - 1)}}{2\beta}$$

Finally the inflection points on the x axis can be obtained from:

$$t = \sqrt[\beta]{\frac{3(\beta - 1) \pm \sqrt{(5\beta - 1)(\beta - 1)}}{2\beta}} \eta \quad (\text{A - 8})$$

It is clear that the inflection points exist only when $\beta > 1$. When $\beta = 2$, one of the inflection points is always 0.

The k^{th} moment about the origin is defined as:

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$$E(x^k) = \int_0^{\infty} t^k f(t) dt$$

For the Weibull distribution, the general expression of the k^{th} moment about the origin is given by [37].

$$E(x^k) = \eta^k \Gamma\left(1 + \frac{k}{\beta}\right)$$

Let $\gamma_k = \Gamma\left(1 + \frac{k}{\beta}\right), k = 0, 1, 2, 3, 4, \dots$

then $E(x^k) = \eta^k \gamma_k$

The first moment can be easily seen as $E(x^k) = E(x) = \text{Mean}$ if $k = 1$, hence,

$$E(x) = \text{Mean} = \eta \gamma_1 = \eta \Gamma\left(1 + \frac{1}{\beta}\right). \quad (\text{A - 9})$$

The variance of the Weibull distribution is defined as:

$$\text{VAR} = E(x^2) - [E(x)]^2$$

This is the second moment minus the square of the first moment.

The second moment is:

$$E(x^2) = \eta^2 \gamma_2 = \eta^2 \Gamma\left(1 + \frac{2}{\beta}\right)$$

Therefore the variance is obtained:

$$\begin{aligned} VAR &= \eta^2 \gamma_2 - (\eta \gamma_1)^2 = \eta^2 (\gamma_2 - \gamma_1^2) \\ &= \eta^2 \left\{ \Gamma\left(1 + \frac{2}{\beta}\right) - \left[\Gamma\left(1 + \frac{1}{\beta}\right) \right]^2 \right\} \end{aligned} \quad (A-10)$$

The standard deviation (SD), therefore, is:

$$SD = \eta \sqrt{\gamma_2 - \gamma_1^2} \quad (A-11)$$

The coefficient of variation (CV) is defined as SD/Mean, thus,

$$CV = \frac{\eta \sqrt{\gamma_2 - \gamma_1^2}}{\eta \gamma_1} = \sqrt{\frac{\gamma_2}{\gamma_1^2} - 1} \quad (A-12)$$

The k^{th} moment about the mean is defined as:

$$E[x - E(x)]^k$$

For the Weibull distribution, the general expression of the k^{th} moment about the mean is given by [37].

$$\begin{aligned} E[x - E(x)]^k &= \eta^k \sum_{i=0}^k (-1)^i \binom{k}{i} \Gamma\left(1 + \frac{k-i}{\beta}\right) \left[\Gamma\left(1 + \frac{1}{\beta}\right) \right]^i \\ &= \eta^k \sum_{i=0}^k (-1)^i \binom{k}{i} \gamma_{k-i} \gamma_1^i \\ &= \eta^k \sum_{i=0}^k (-1)^i \frac{k!}{(k-i)! i!} \gamma_{k-i} \gamma_1^i \end{aligned}$$

The skewness is defined as the expected value of the third moment about the mean $E[x - E(x)]^3$.

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The coefficient of skewness (CS) is defined as $\frac{E[x - E(x)]^3}{VAR^{\frac{3}{2}}}$:

$$\begin{aligned}
 CS &= \frac{\eta^3 \left(\frac{3!}{3!0!} \gamma_3 \gamma_1^0 - \frac{3!}{2!1!} \gamma_2 \gamma_1^1 + \frac{3!}{1!2!} \gamma_1 \gamma_1^2 - \frac{3!}{0!3!} \gamma_0 \gamma_1^3 \right)}{\left[\eta^2 (\gamma_2 - \gamma_1^2) \right]^{\frac{3}{2}}} \\
 &= \frac{\gamma_3 - 3\gamma_2 \gamma_1 + 2\gamma_1^3}{\left[(\gamma_2 - \gamma_1^2) \right]^{\frac{3}{2}}} \tag{A-13}
 \end{aligned}$$

The kurtosis is defined as the expected value of the fourth moment about the mean $E[x - E(x)]^4$.

The coefficient of kurtosis(CK) is defined as $\frac{E[x - E(x)]^4}{VAR^{\frac{4}{2}}}$:

$$\begin{aligned}
 CK &= \frac{\eta^4 \left(\frac{4!}{4!0!} \gamma_4 \gamma_1^0 - \frac{4!}{3!1!} \gamma_3 \gamma_1^1 + \frac{4!}{2!2!} \gamma_2 \gamma_1^2 - \frac{4!}{1!3!} \gamma_1 \gamma_1^3 + \frac{4!}{0!4!} \gamma_0 \gamma_1^4 \right)}{\left[\eta^2 (\gamma_2 - \gamma_1^2) \right]^2} \\
 &= \frac{\gamma_4 - 4\gamma_3 \gamma_1 + 6\gamma_2 \gamma_1^2 - 3\gamma_1^4}{(\gamma_2 - \gamma_1^2)^2} \tag{A-14}
 \end{aligned}$$